

Sharp bounds for transition probability densities of a class of diffusions

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Abstract Let $p_b(x, t, y)$ be the transition probability density of the one dimensional diffusion process $dX_t = dW_t + b(X_t) dt$, where $|b(\cdot)|_\infty \leq 1$. We show that the upper and lower bounds of $p_b(x, t, y)$ are achieved for fixed (x, t, y) when $b(z) = \text{sgn}(y - z)$ and $b(z) = \text{sgn}(z - y)$ respectively. Moreover, the precise bounds are given. *To cite this article: Z. Qian, W. Zheng, C. R. Acad. Sci. Paris, Ser. I 335 (2002) 953–957.*

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Bornes explicites pour la densité de la probabilité de transition d'un processus de diffusion

Résumé Soit $p_b(x, t, y)$ la densité de la probabilité de transition du processus de diffusion $dX_t = dW_t + b(X_t) dt$, où $|b(\cdot)|_\infty \leq 1$. Nous montrons que la borne supérieure (resp. la borne inférieure) de $p_b(x, t, y)$ est atteinte pour (x, t, y) fixés quand $b(z) = \text{sgn}(y - z)$ (resp. quand $b(z) = \text{sgn}(z - y)$). Les bornes explicites sont présentées. *Pour citer cet article : Z. Qian, W. Zheng, C. R. Acad. Sci. Paris, Ser. I 335 (2002) 953–957.*

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Considérons l'équation différentielle stochastique

$$dX_t = b(X_t) dt + dW_t, \quad (1)$$

où b est une fonction mesurable bornée sur \mathbb{R} , et $(W_t)_{t \geq 0}$ est le mouvement brownien standard. Dans cette Note, nous prouvons que la densité de la probabilité de transition $p_b(x, t, y)$ de (X_t) vérifie les bornes suivantes.

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THÉORÈME 1. – Soit $\beta = |b|_\infty < \infty$. Alors pour tous x, y et $t > 0$

$$\frac{1}{\sqrt{2\pi t}} \int_{|x-y|/\sqrt{t}}^{\infty} z e^{-(z+\beta\sqrt{t})^2/2} dz \leq p_b(x, t, y) \leq \frac{1}{\sqrt{2\pi t}} \int_{|x-y|/\sqrt{t}}^{\infty} z e^{-(z-\beta\sqrt{t})^2/2} dz. \quad (2)$$

De plus, nous avons le

THÉORÈME 2. – Soit $y \in \mathbb{R}$ et β une constante. Notons par $p_y^\beta(x, t, z)$ la densité de la probabilité (par rapport à la mesure de Lebesgue) associée au processus de diffusion

$$dX_t = \beta \operatorname{sgn}(y - X_t) dt + dW_t. \quad (3)$$

Alors

$$p_y^\beta(x, t, y) = \frac{1}{\sqrt{2\pi t}} \int_{|x-y|/\sqrt{t}}^{\infty} z e^{-(z-\beta\sqrt{t})^2/2} dz. \quad (4)$$

Consider the following stochastic differential equation on a complete probability space $(\Omega, \mathcal{F}, \mathbb{P})$:

$$dX_t = b(X_t) dt + dW_t, \quad (5)$$

where b is a bounded measurable function on \mathbb{R} , and $(W_t)_{t \geq 0}$ is the standard Brownian motion. It is well known that there is a unique weak solution to (5) and its transition probability density (with respect to the Lebesgue measure) $p_b(x, t, y)$ is Hölder continuous in all variables (see, e.g., [1,6]). In this Note we provide the optimal upper and lower bounds for the transition probability density $p_b(x, t, y)$ in terms of the bound $|b|_\infty$. Let us state our main result as the following:

THEOREM 1. – Assume $\beta = |b|_\infty < \infty$. Then for all x, y and $t > 0$

$$\frac{1}{\sqrt{2\pi t}} \int_{|x-y|/\sqrt{t}}^{\infty} z e^{-(z+\beta\sqrt{t})^2/2} dz \leq p_b(x, t, y) \leq \frac{1}{\sqrt{2\pi t}} \int_{|x-y|/\sqrt{t}}^{\infty} z e^{-(z-\beta\sqrt{t})^2/2} dz. \quad (6)$$

The bounds in (6) are optimal under the norm $|b|_\infty$, see Theorem 2. Our proof depends on the precise information on the transition probability densities for the extremal vector fields b which we state as the following theorem, which has independent interest on its own.

THEOREM 2. – Let $y \in \mathbb{R}$ and let β be a constant. Denote by $p_y^\beta(x, t, z)$ the transition probability density (with respect to the Lebesgue measure) associated to the diffusion process

$$dX_t = \beta \operatorname{sgn}(y - X_t) dt + dW_t. \quad (7)$$

Then

$$p_y^\beta(x, t, y) = \frac{1}{\sqrt{2\pi t}} \int_{|x-y|/\sqrt{t}}^{\infty} z e^{-(z-\beta\sqrt{t})^2/2} dz. \quad (8)$$

Proof. – It is sufficient to consider the case where $|x - y| > 0$. The case where $|x - y| = 0$ follows the continuity of both sides of (8). Note that $p_y^\beta(x, t, z)$ is continuous in (x, t, z) . Therefore

$$\frac{d}{d\lambda} \mathbb{P}(|X_t - y| \leq \lambda) = p_y^\beta(x, t, y + \lambda) + p_y^\beta(x, t, y - \lambda)$$

for all $\lambda \geq 0$. In particular

$$p_y^\beta(x, t, y) = \frac{1}{2} \frac{d}{d\lambda} \Big|_{\lambda=0+} \mathbb{P}(|X_t - y| \leq \lambda).$$

Since

$$dX_t = \beta \operatorname{sgn}(y - X_t) dt + dW_t; \quad X_0 = x,$$

by Tanaka's formula ([3], Proposition 3.6.8 or [5], Theorem VI.1.2)

$$\begin{aligned} |X_t - y| &= |x - y| + \int_0^t \operatorname{sgn}(X_s - y) dX_s + L_t \\ &= |x - y| - \beta t + \int_0^t \operatorname{sgn}(X_s - y) dW_s + L_t, \end{aligned}$$

where $(L_t)_{t \geq 0}$ is the local time of the process $(X_t)_{t \geq 0}$ at y . The local time L_t is given by the Skorohod equation ([3], Lemma 3.6.14)

$$L_t = \max \left\{ 0, \sup_{s \leq t} \left(-|x - y| + \beta s - \int_0^s \operatorname{sgn}(X_u - y) dW_u \right) \right\}.$$

Define $B_t = \beta t - \int_0^t \operatorname{sgn}(X_s - y) dW_s$, and define the probability measure \mathbb{Q} by

$$\frac{d\mathbb{Q}}{d\mathbb{P}} \Big|_{\mathcal{F}_t} = \exp \left(\int_0^t \beta \operatorname{sgn}(X_s - y) dW_s - \frac{\beta^2 t}{2} \right).$$

Then by the Cameron–Martin formula, under \mathbb{Q} , $(B_t)_{t \geq 0}$ is a standard Brownian motion with initial value zero. Let $M_t = \sup_{s \leq t} B_s$ be the maximum process of B_t . Then we have

$$|X_t - y| = |x - y| - B_t + L_t$$

and

$$L_t = \max \{ 0, -|x - y| + M_t \},$$

so that

$$|X_t - y| = -B_t + \max \{ |x - y|, M_t \}.$$

Since (B_t, M_t) possesses the joint distribution

$$\mathbb{Q}(B_t \in da, M_t \in db) = \frac{2(2b - a)}{\sqrt{2\pi t^3}} \exp \left\{ -\frac{(2b - a)^2}{(2t)} \right\} da db$$

on $a \leq b$ and $b \geq 0$ (see [3], formula (2.8.2)), therefore when $0 < \lambda < |x - y|$

$$\begin{aligned} \mathbb{P}(|X_t - y| \leq \lambda) &= \mathbb{Q} \left(\exp \left(\beta B_t - \frac{\beta^2 t}{2} \right) : |X_t - y| \leq \lambda \right) \\ &= \mathbb{Q} \left(\exp \left(\beta B_t - \frac{\beta^2 t}{2} \right) : -B_t + \max \{ |x - y|, M_t \} \leq \lambda \right) \\ &= \iint_{\substack{a \leq b; b \geq 0 \\ -a + \max \{ |x - y|, b \} \leq \lambda}} e^{\beta a - \beta^2 t / 2} \frac{2(2b - a)}{\sqrt{2\pi t^3}} e^{-(2b - a)^2 / (2t)} da db \end{aligned}$$

$$\begin{aligned} &= \int_{|x-y|-\lambda}^{+\infty} \int_a^{a+\lambda} e^{\beta a - \beta^2 t/2} \frac{2(2b-a)}{\sqrt{2\pi t^3}} e^{-(2b-a)^2/(2t)} db da \\ &= \frac{1}{\sqrt{2\pi t}} \int_{|x-y|-\lambda}^{+\infty} e^{\beta a - \beta^2 t/2} (e^{-a^2/(2t)} - e^{-(a+2\lambda)^2/(2t)}) da. \end{aligned}$$

Taking derivative in λ we thus obtain

$$\begin{aligned} \frac{d}{d\lambda} \mathbb{P}(|X_t - y| \leq \lambda) &= \frac{1}{\sqrt{2\pi t}} e^{\beta(|x-y|-\lambda)-\beta^2 t/2} (e^{-(|x-y|-\lambda)^2/(2t)} - e^{-(|x-y|+\lambda)^2/(2t)}) \\ &\quad + \frac{e^{-\beta^2 t/2}}{\sqrt{2\pi t}} \int_{|x-y|-\lambda}^{+\infty} e^{-(a+2\lambda)^2/(2t)+\beta a} \frac{2(a+2\lambda)}{t} da. \end{aligned}$$

Letting $\lambda \downarrow 0$ we hence establish

$$p_y^\beta(x, t, y) = \frac{e^{-\beta^2 t/2}}{\sqrt{2\pi t}} \int_{|x-y|/\sqrt{t}}^{+\infty} z e^{-z^2/2+\sqrt{t}\beta z} dz. \quad \square$$

Now we are in a position to prove Theorem 1. Without losing generality we may assume that $|b|_\infty = 1$, and use $p_y^+(x, t, z)$ (resp. $p_y^-(x, t, z)$) to denote $p_y^\beta(x, t, z)$ when $\beta = 1$ (resp. when $\beta = -1$). Since

$$\frac{\partial}{\partial x} p_y^\beta(x, t, y) = -\frac{x-y}{t\sqrt{2\pi t}} e^{-(x-y)^2/(2t)-\beta^2 t/2+\beta|x-y|},$$

therefore

$$(b(x) - \text{sgn}(y-x)) \frac{\partial}{\partial x} p_y^+(x, t, y) \leq 0$$

as $|b|_\infty \leq 1$. Let X_t be the diffusion process:

$$dX_t = \text{sgn}(y - X_t) dt + dW_t$$

and $\mathbb{E}_{x,y}$ denote the probability of the process $(X_s)_{0 \leq s \leq t}$ conditional on $X_0 = x$ and $X_t = y$. By the Cameron-Martin formula

$$\begin{aligned} \frac{p_b(x, t, y)}{p_y^+(x, t, y)} &= \mathbb{E}_{x,y} \exp \left\{ \int_0^t (b(X_s) - \text{sgn}(y - X_s)) dW_s - \frac{1}{2} \int_0^t (b(X_s) - \text{sgn}(y - X_s))^2 ds \right\} \\ &= \mathbb{E}_{x,y} \exp \left\{ \int_0^t (b(X_s) - \text{sgn}(y - X_s)) d\tilde{W}_s - \frac{1}{2} \int_0^t (b(X_s) - \text{sgn}(y - X_s))^2 ds \right. \\ &\quad \left. + \int_0^t (b(X_s) - \text{sgn}(y - X_s)) \frac{\partial}{\partial x} \log p_y^+(X_s, t-s, y) ds \right\} \\ &\leq \mathbb{E}_{x,y} \exp \left\{ \int_0^t (b(X_s) - \text{sgn}(y - X_s)) d\tilde{W}_s - \frac{1}{2} \int_0^t (b(X_s) - \text{sgn}(y - X_s))^2 ds \right\} \\ &= 1, \end{aligned}$$

where \tilde{W} is a standard Brownian motion under the conditional probability $\mathbb{P}_{x,y}$ (see [4]) which yields that

$$\frac{p_b(x, t, y)}{p_y^+(x, t, y)} \leq 1.$$

Thus we have proved the upper bound in (6). Similarly, it is easy to see that

$$(b(x) + \operatorname{sgn}(y - x)) \frac{\partial}{\partial x} p_y^-(x, t, y) \geq 0$$

which in turn proves the lower bound in (6).

After we submitted the paper, we have learned that a quite similar formula for the density $p_y^+(x, t, z)$ has been obtained in [2] by Grădinaru, Herrmann and Roynette. In fact, in [2] a formula for $p_y^+(y, t, x)$ is given, while ours for $p_y^+(x, t, y)$, but they seem equivalent. Also we notice that there is a formula (but a little bit complicated) in ([3], formula (6.5.14)) for the density $p_y^+(x, t, z)$ under the name of diffusions with two-valued drifts. However we believe that our formula (8) is the most clean one, and our contribution thus is to bring these facts into a comparison type theorem (see Theorem 1) which we believe is useful in applications.

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